



## LIBBY-NOVICK KUMARASWAMY DISTRIBUTION WITH ITS PROPERTIES AND APPLICATIONS

ABDUL SABOOR<sup>1,\*</sup>, ZAFAR IQBAL<sup>2</sup>, MUHAMMAD HANIF<sup>1</sup>, MUNIR AHMAD<sup>1</sup>

<sup>1</sup>*School of social sciences, National College of Business Administration & Economics, Lahore, Pakistan*

<sup>2</sup>*Department of Statistics, Government College Satellite Town Gujranwala, Pakistan*

*\*Corresponding author: [saboorsd@hotmail.com](mailto:saboorsd@hotmail.com)*

ABSTRACT: The Kumaraswamy distribution is one of the most popular probability distributions with applications to real life data. In this paper, an extension of this distribution called the Libby-Novick Kumaraswamy (LNK) distribution is presented which is believed to provide greater flexibility to model scenarios involving skew-normal data than original one. Analytical expressions for various mathematical properties including its cdf, quantile function, moments, factorial moments, conditional moments, moment generating function, characteristic function, vitality function, information generating function, reliability measures, mean deviations, mean residual function, Bonferroni and Lorenz Curves are derived. The parameters' estimation of LNK distribution is undertaken using the method of maximum likelihood estimation. A simulation study for different values of sample sizes, to assess the performance of the parameters of LNK distribution is provided. For illustration and performance evaluation of LNK distribution three real-life data sets from the field of engineering and science adapted from earlier studies are used. On comparing the results to previously used methods, LNK distribution shows that it can give consistently better fit than other existing important lifetime models. It is found that the LNK distribution is more suitable and useful to study lifetime data.

---

Received September 1<sup>st</sup>, 2020; accepted September 16<sup>th</sup>, 2020; published April 12<sup>th</sup>, 2021.

2010 Mathematics Subject Classification. 62E17, 05A15.

Key words and phrases. Libby-Novick Kumaraswamy distribution; moment generating function; mean deviation; reliability measures; maximum likelihood estimation.

©2021 Authors retain the copyrights

of their papers, and all open access articles are distributed under the terms of the Creative Commons Attribution License.

## 1. INTRODUCTION

Iqbal et al., (2017) [1] defined the Libby-Novick Kumaraswamy (LNK) distribution. The probability density function (*pdf*) and cumulative distribution function (*cdf*) are respectively defined as

$$f(x) = abc \frac{x^{a-1} (1-x^a)^{b-1}}{[1-(1-c)x^a]^{1+b}} ; \quad 0 < x < 1, \quad a, b, c > 0 \quad (1)$$

$$F(x) = 1 - \left( \frac{1-x^a}{1-(1-c)x^a} \right)^b \quad (2)$$

The Libby-Novick Kumaraswamy (LNK) distribution is a continuous probability distribution with double-bounded support. It is very similar, in many respects, to the Libby-Novick Beta (LNB, 1982) [2] distribution and Kumaraswamy (Kum, 1980) [3] distribution. The Kum distribution is a two parameter distribution like Beta distribution Where as M.C. Jones (2009) [4] find out numerous benefits of the Kum distribution over classical beta distribution. The Kum distribution is a special case of McDonald's (1984) [5] generalized Beta of the first kind distribution. One key difference between the Kum and Beta distributions is the availability for the former, but not for the latter, of an invertible closed form cumulative distribution function presented by Mitnick (2013) [6].

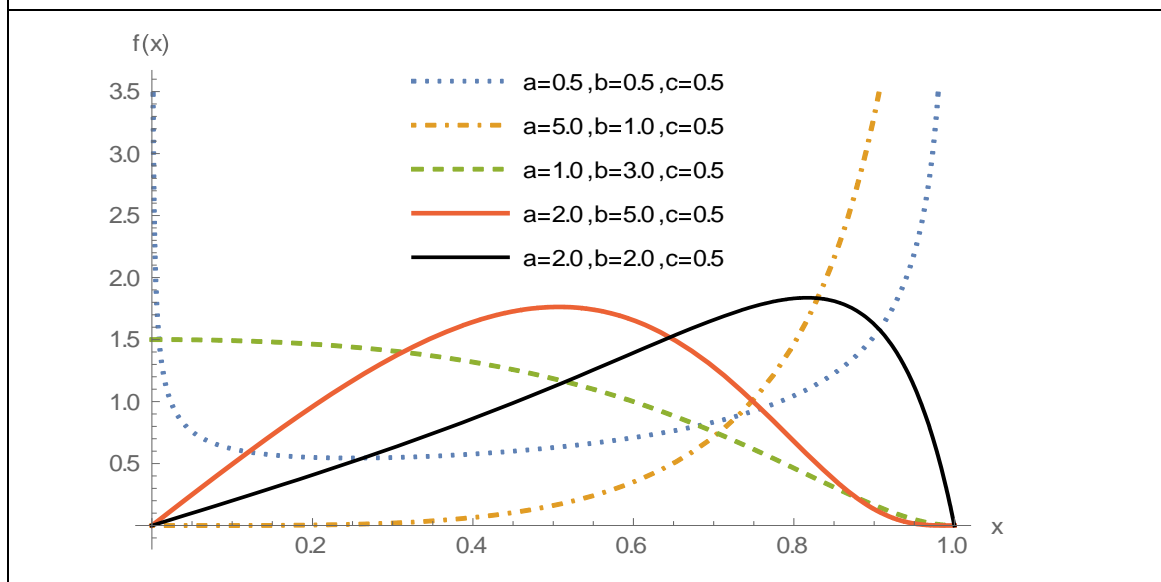
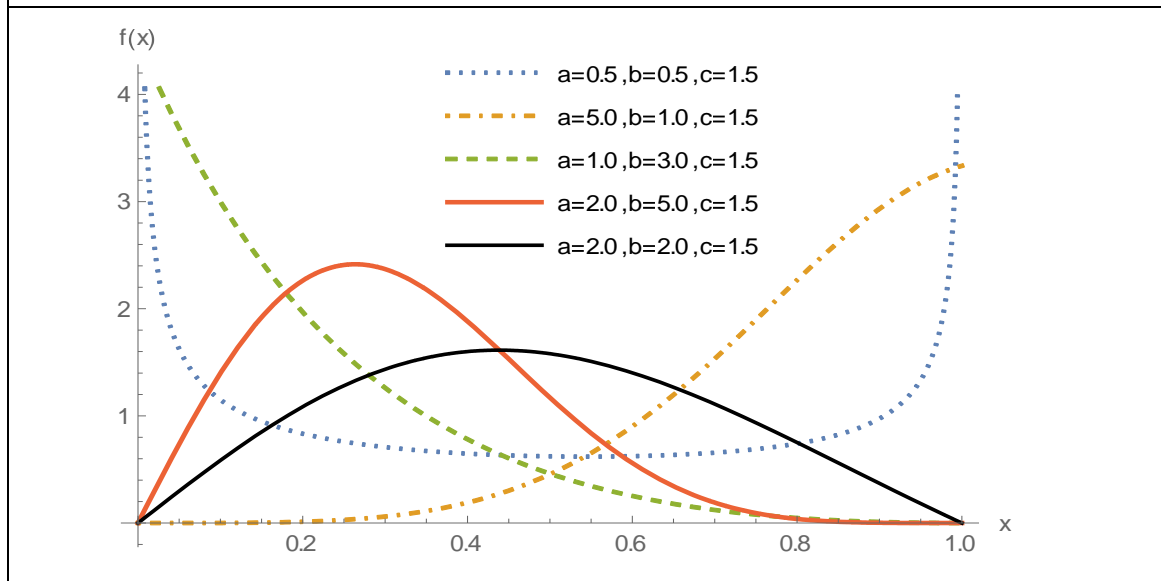
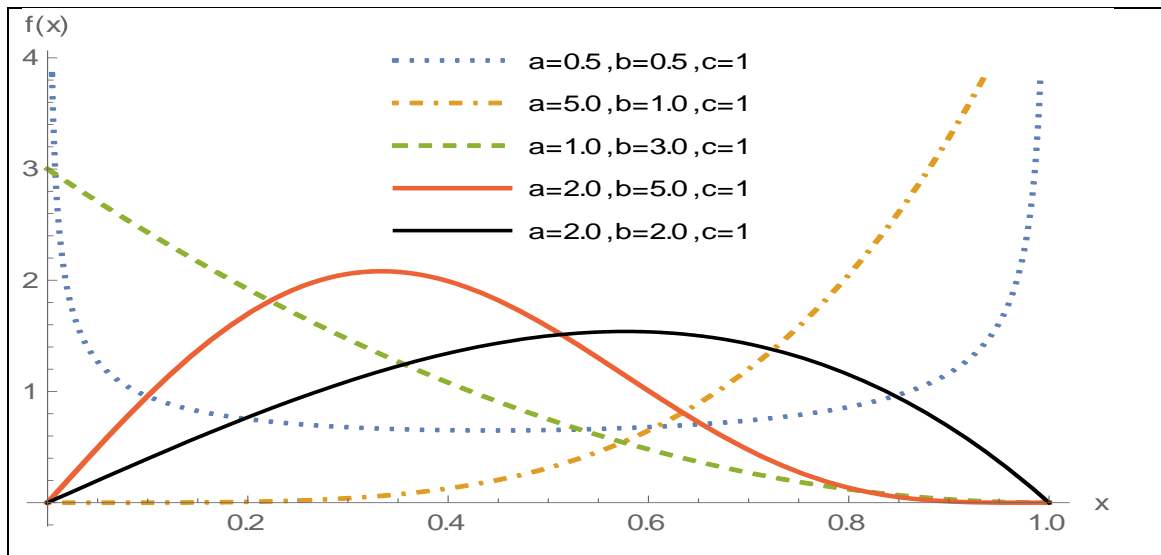
The two distributions LNK and LNB are very flexible and can take approximately the same shapes and the former distribution is also a generalized distribution of Kum distribution. There are, however, important realistic differences between LNB and LNK distributions. On the one hand, the availability for the LNK, but not for the LNB distribution, of an invertible closed-form cumulative distribution function makes the LNK distribution much better suited than the LNB for activities that require the generation of random varieties, in particular simulation modeling and simulation-based model estimation. In contrast, the lack of tractable-enough expressions for the mean and variance of the LNK distribution has stalled its utilization for modeling purposes; in spite of the advantages that the availability of an invertible closed-form cumulative distribution function entails, the LNK distribution may use rather sparingly in the modeling of stochastic phenomena and processes. Libby-Novick (1982) [2] derived two new multivariate probability density functions, which are the generalized forms of beta distribution

and F distribution. They proved that these distributions that seem to be suitable in utility modeling. They reproduced the same distributional form in the cases of marginal and conditional distributions. Chen and Novick (1984) [7] used the Libby-Novick (1982) [2] univariate generalized beta distribution as a prior and estimated the parameters of Bernoulli and Binomial distributions. These expressions are the generalized forms of the standard beta class. Ristic et al., (2013) [8] derived new family of skewed distributions such as Libby and Novick's generalized beta exponential distribution and found some useful properties of this family of distributions. Cordeiro et al., (2014) [9] defined a family of distributions, named the Libby-Novick beta family of distributions, which includes the classical beta generalized and exponentiated generators. This extended family gave reasonable parametric fits to real data in several areas because the additional shape parameters controlled the skewness and kurtosis simultaneously. Ali (2019) [10] worked on new form of Libby-Novick (NLN) distribution and explored some properties of NLN distribution. This model was compared with other distributions by fitting them to a real data set. Rashid et al., (2020) [11] derived different entropy measures and characterized Libby-Novick generalized Beta (LNGB) distribution through various methods. Iqbal et al., (2021) [12] derived mathematical properties of LNGB distribution and applied it to modeling on three real data sets. In this paper we derive basic and advanced properties of LNK distribution and find applications of LNK distribution to three real data sets. In Section 2, a detailed remarks about the graphs of pdf, hazard rate function, reverse hazard rate function and survival function are provided. Section 2 also contains the derivation and results of some important mathematical properties of LNK distribution. Maximum likelihood estimators of three parameters of LNK distribution are derived in section 3. In section 4, a simulation of the parameters is carried out for different sample sizes. A number of deduced models are shown in section 5. In section 6, LNK distribution is compared with some other models through three data sets. Some concluding remarks are presented in section 7.

## 2. MATHEMATICAL PROPERTIES

### 1.1 Shape properties of the *pdf*

The LNK density function defined in (1) has real flexibility and it is shown through graphs w.r.t. some different combinations of values of the parameters.



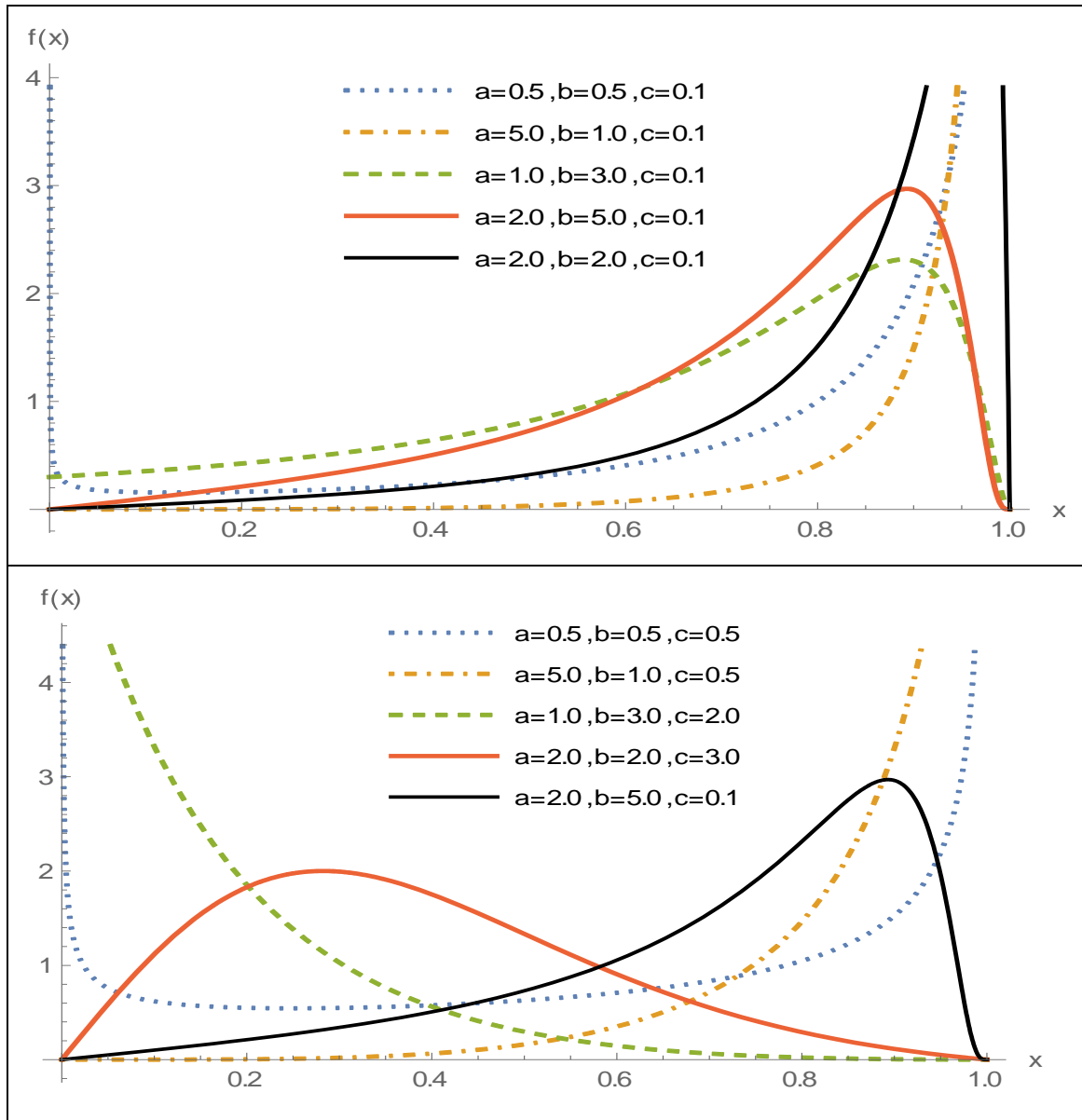


Figure 2.1 pdf Graph of LNK distribution

**Shapes of pdf**

- (i) For  $a > 0, c > 0$  and  $0 < b < 1$  then LNK distribution is U-shaped.
- (ii) For  $0 < a < 1, b > 1$  and  $0 < c < 1$  then LNK distribution is S-shaped.
- (iii) For  $a = b = 2, c \geq 1$  ( $c < 1$ ) and if  $c \rightarrow \infty$  ( $c \rightarrow 0$ ) then LNK distribution increases its positively skewed (negatively skewed) from the symmetric with decreasing mode (increasing mode).
- (iv) For  $a = b = c = 1$ , the LNK distribution is uniform distribution.
- (v) For  $a > 2, b > 2$  and  $c \geq 1$  the LNK distribution is unimodal positively skewed with decreasing mode when  $c \rightarrow \infty$ .
- (vi) For  $a > 2, b = 1, 0 < c < 1$ , the form of LNK distribution is an increasing.

- (vii) For  $a > 2, b = 1, 1 < c < 3.5$ , the LNK distribution increase but it increases slowly when  $c$  increases in the interval and for  $c \geq 3.5$  the LNK distribution again turns to unimodel.
- (viii) For  $b = c = 1$ , the LNK distribution is Power distribution.
- (ix) For  $a = c = 1$ , the LNK distribution is LNK is a special case of Kum-distribution or reflected exponentiated distribution.

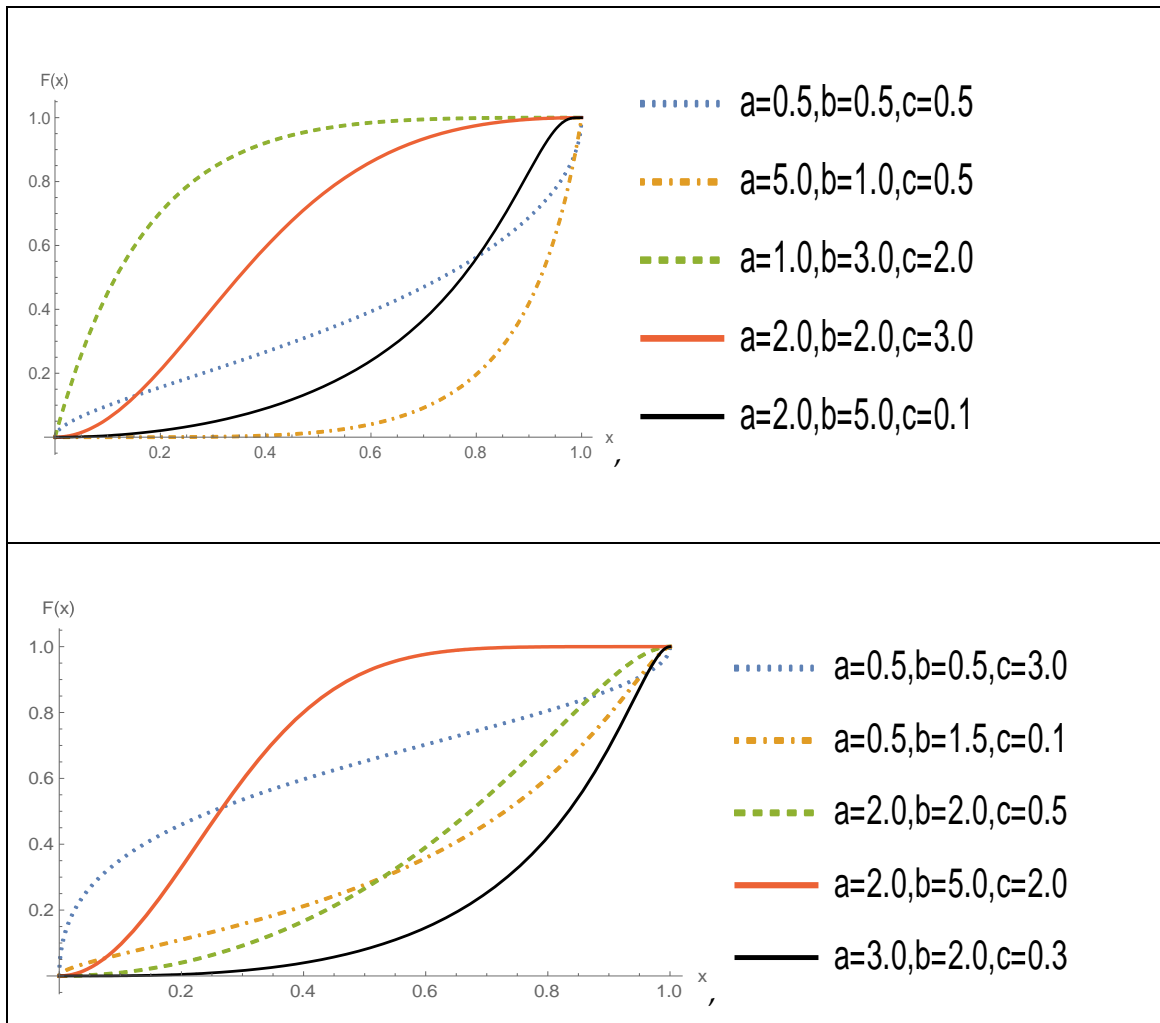
**1.2 Distribution function**

The cumulative distribution function (*cdf*) of LNK distribution is

$$F(x) = abc \int_0^x \frac{t^{a-1} (1-t^a)^{b-1}}{[1-(1-c)t^a]^{1+b}} dt \tag{3}$$

By making substitutions  $t^a = z$  and  $u = \frac{cz}{1-z+cz}$  (4)

And after some simplification we have the expression of *cdf* as shown in (2)



**Figure 2.2** *cdf* Graph of LNK distribution

### 2.3 Quantile Function

The quantile function of LNK distribution is given by:

$$F(x) = p \quad \text{Where } 0 < p < 1$$

And it  $x = F^{-1}(p)$

$$1 - \left[ \frac{1 - x^a}{1 - (1 - c)x^a} \right]^b = p$$

After simplification we find the quantile function of the LNK distribution as under;

$$Q(p; a, b, c) = \left[ \frac{1 - (1 - p)^{1/b}}{1 - (1 - c)(1 - p)^{1/b}} \right]^{1/a} \quad p \in (0, 1)$$

Where  $Q(p; a, b, c)$  is the inverse of LNK function or quantile of LNK function at  $p$ . Clearly, the function has explicit form and it can be numerically solved through software for different set of parameters' values. The graph can also be used to illustrate the behaviour of quantile function of LNK distribution.

- i. For  $a > 1, b > 1$  by increasing  $c$ , the quantile value of the LNK distribution decreases comparatively.
- ii. For  $a < 1, c > 1$ , and for any value of  $b$ , the quantile value of the LNK distribution increases slowly for  $p < 0.5$  and for  $p > 0.5$ , the quantile value sharply increases.
- iii. For  $c \rightarrow 0$ , the quantile value of the LNK distribution increases sharply.

### 2.4 Hazard Rate Function

The hazard or instantaneous rate function is denoted by  $h(x)$ . The hazard function of  $x$  can be interpreted as instantaneous rate or the conditional probability density of failure at time  $x$ , given that the unit has survived until  $x$ . The hazard rate function of LNK distribution as

$$h(x; a, b, c) = \frac{abcx^{a-1}}{\left[ 1 - (1 - c)x^a \right] (1 - x^a)}$$

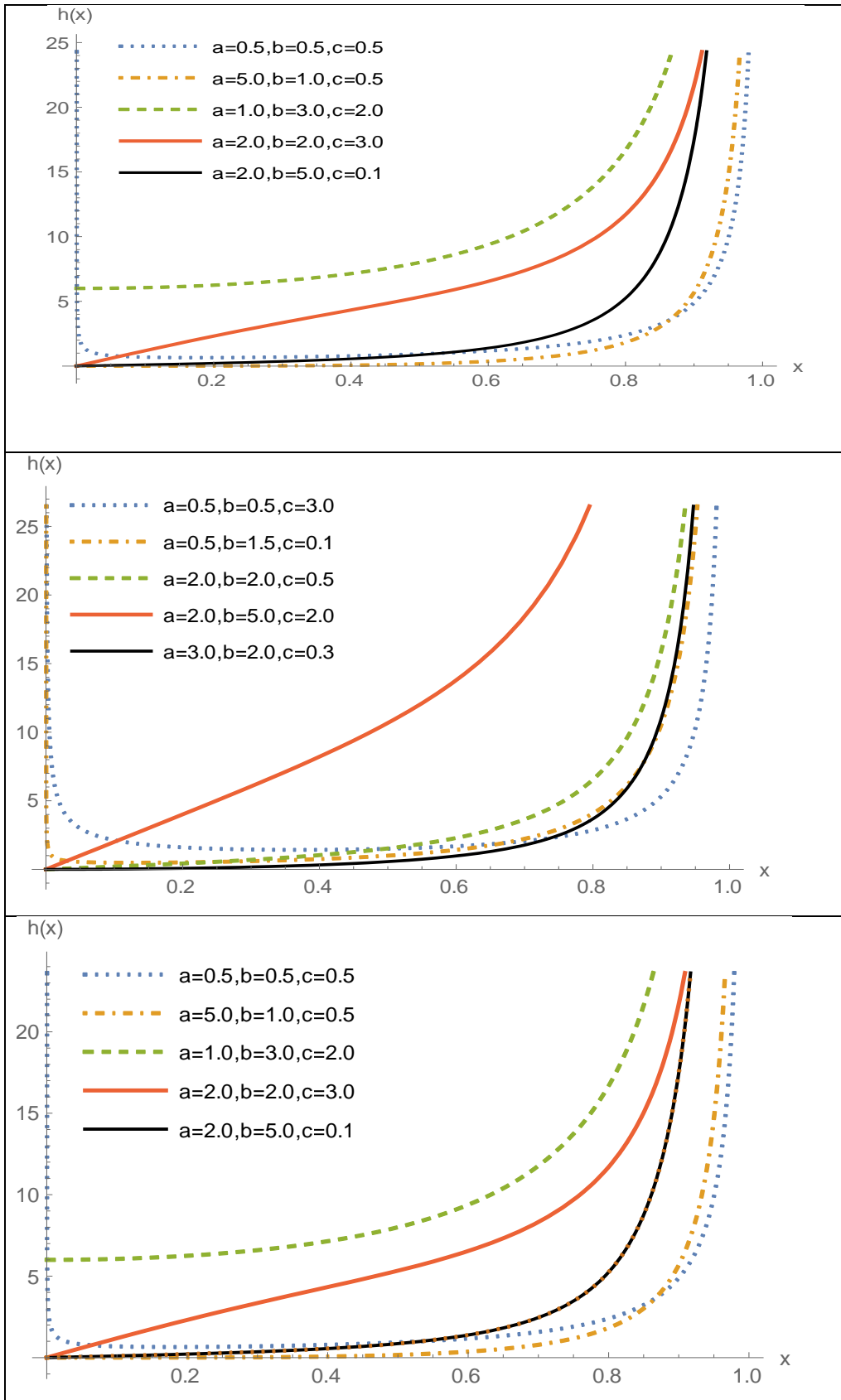


Figure 2.3 hazard rate function Graph of LNK distribution

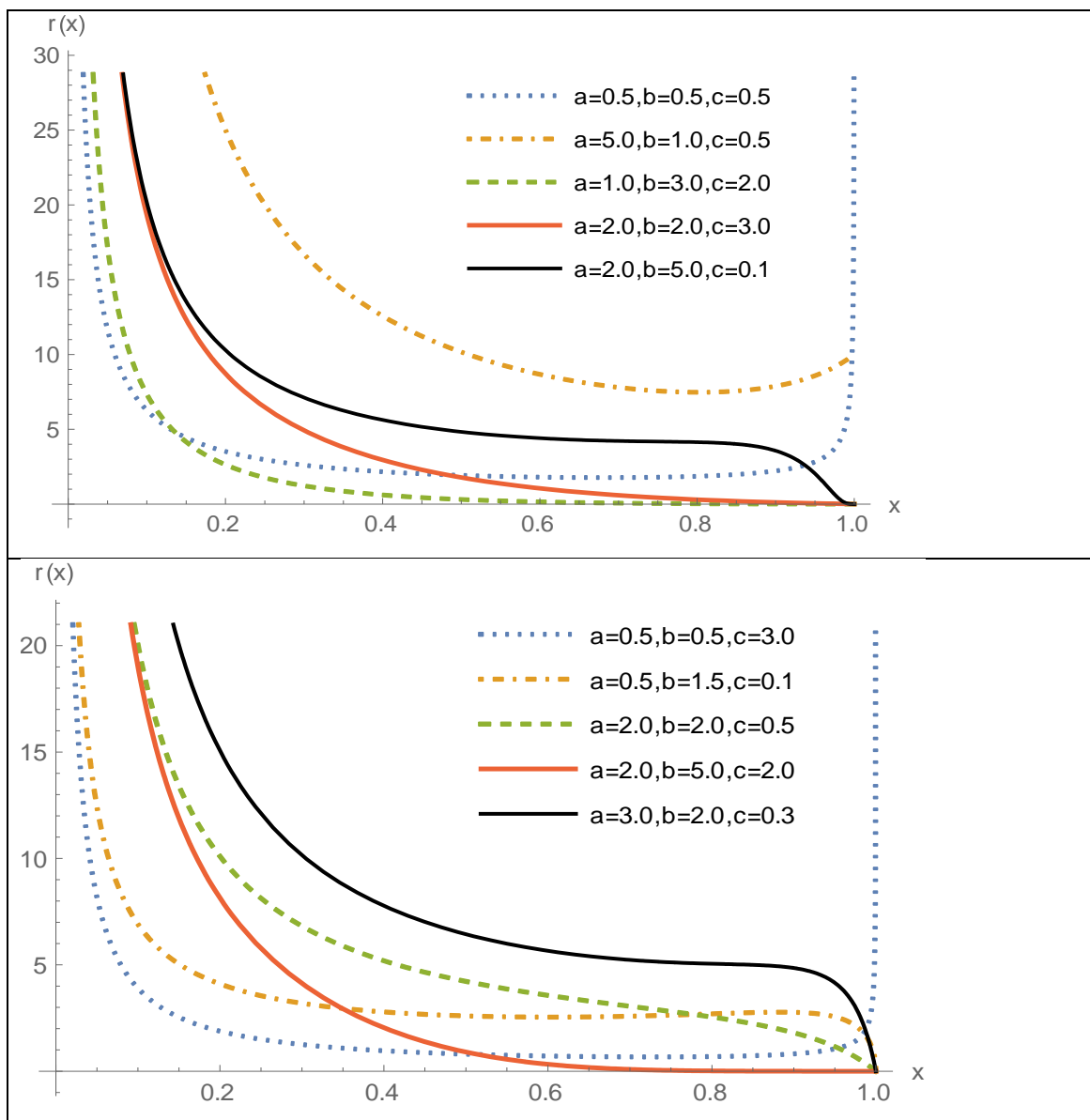


The hazard rate function of LNK distribution is of Bath-Tub shape, increasing shape and decreasing shape for specific sets of values of parameters.

**2.5 Reverse Hazard rate function**

The reverse hazard can be interpreted as an approximate probability of failure in  $[x, x + d]$ , given that the failure has occurred in  $[0, x]$ . The reverse hazard function  $r(x; a, b, c)$  of LNK distribution is defined as

$$r(x; a, b, c) = \frac{abcx^{a-1}(1-x^a)^{b-1}}{(1-(1-c)x^a)\left(\left(1-(1-c)x^a\right)^b - (1-x^a)^b\right)}$$



**Figure 2.4** reverse hazard rate Graph of LNK distribution

### 2.6 Survival Function

The survival function or reliability function of LNK distribution is defined as

$$S(x) = \left[ \frac{1 - x^a}{1 - (1 - c)x^a} \right]^b$$

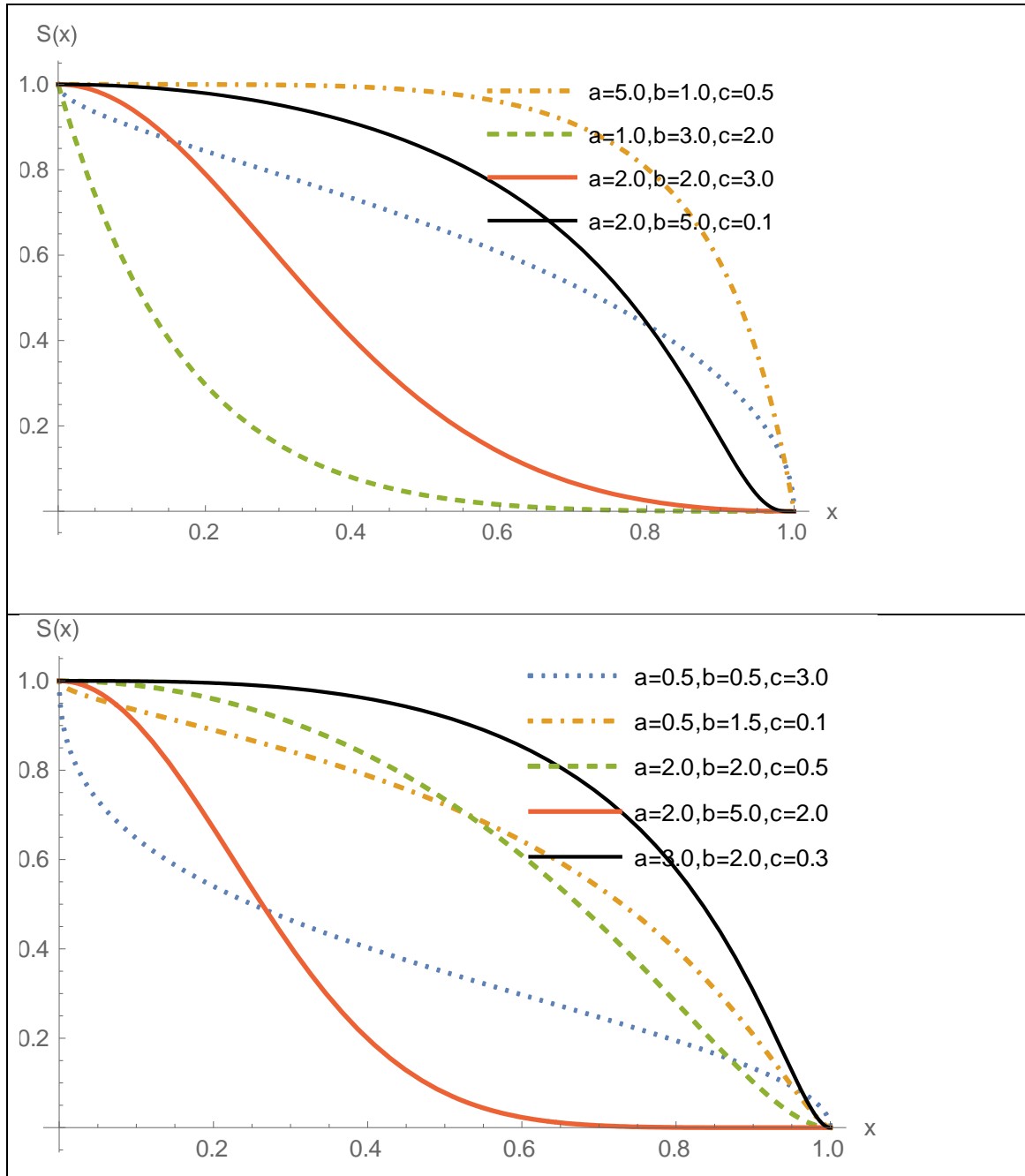


Figure 2.5 survival function Graph of LNK distribution

## 2.7 The $r^{\text{th}}$ Moment

The  $r^{\text{th}}$  moments about origin of the LNK distribution is defined as

$$E(x^r) = \int_0^1 x^r f(x) dx$$

By Substitution (4) becomes

$$\begin{aligned} &= \int_0^1 \frac{bcy^{\frac{r}{a}}(1-y)^{b-1}}{[1-(1-c)y]^{1+b}} dy \\ &= b \int_0^1 \left( \frac{u}{c+u-uc} \right)^{\frac{r}{a}} (1-u)^{b-1} du \\ &= bc^{\frac{-r}{a}} \int_0^1 u^{\frac{r}{a}} (1-u)^{b-1} \sum_{k=0}^{\infty} \frac{\Gamma(r/a+k)}{\Gamma(r/a)k!} \left(1-\frac{1}{c}\right)^k u^k du \\ &= bc^{\frac{-r}{a}} \sum_{k=0}^{\infty} \frac{\Gamma(r/a+k)}{\Gamma(r/a)k!} \left(1-\frac{1}{c}\right)^k \int_0^1 u^{\frac{r}{a}+k} (1-u)^{b-1} du \\ E(x^r) &= \frac{b}{c^{r/a}} \sum_{k=0}^{\infty} \frac{\Gamma(r/a+k)}{\Gamma(r/a)k!} \left(1-\frac{1}{c}\right)^k B(r/a+k+1, b) \end{aligned}$$

**Corollary**

$$\begin{aligned} E(x) &= \frac{b}{c^{1/a}} \sum_{k=0}^{\infty} \frac{\Gamma(1/a+k)}{\Gamma(1/a)k!} \left(1-\frac{1}{c}\right)^k B(1/a+k+1, b) \\ E(x^2) &= \frac{b}{c^{2/a}} \sum_{k=0}^{\infty} \frac{\Gamma(2/a+k)}{\Gamma(2/a)k!} \left(1-\frac{1}{c}\right)^k B(2/a+k+1, b) \end{aligned}$$

## 2.8 Moment generating function

The moment generating function of the LNK distribution about zero is

$$M_x(t) = E(e^{tx})$$

$$= \sum_{r=0}^{\infty} \frac{bc^{\frac{-r}{a}}}{\Gamma\left(\frac{r}{a}\right)} \sum_{k=0}^{\infty} \frac{\Gamma\left(\frac{r}{a}+k\right)}{\Gamma\left(\frac{r}{a}\right)k!} \left(1-\frac{1}{c}\right)^k B\left(\frac{r}{a}+k+1, b\right) \frac{(t^r)}{k!r!} \quad \because e^x = \sum_{r=0}^{\infty} \frac{x^r}{r!}$$

### 2.9 Factorial Moments

This section devotes to increasing and decreasing factorial moments of LNK distribution as under:

#### 2.9.1 Decreasing Factorial Moments of LNK distribution

The decreasing factorial moments of the LNK distribution is defined as:

$$\begin{aligned}
 E[(X)_n] &= E[X(X-1)(X-2) \dots (X-n+1)] \\
 &= \sum_{r=0}^n s(n, r) E(X^r) = \sum_{r=0}^n \sum_{k=0}^{\infty} \frac{s(n, r)}{\Gamma\left(\frac{r}{a}\right) k!} bc^{\frac{-r}{a}} \Gamma\left(\frac{r}{a} + k\right) \left(1 - \frac{1}{c}\right)^k B\left(\frac{r}{a} + k + 1, b\right)
 \end{aligned}$$

Where  $s(n, r)$  are the Stirling's numbers of first kind.

#### 2.9.2 Increasing Factorial Moments of LNBD

The increasing factorial moments of the LNK distribution is defined as:

$$\begin{aligned}
 E[X^{(r)}] &= E[X(X+1)(X+2) \dots (X+r-1)] \\
 &= \sum_{r=0}^n d(n, r) E(X^r) \\
 &= \sum_{r=0}^n \sum_{k=0}^{\infty} \frac{d(n, r)}{\Gamma\left(\frac{r}{a}\right) k!} bc^{\frac{-r}{a}} \Gamma\left(\frac{r}{a} + k\right) \left(1 - \frac{1}{c}\right)^k B\left(\frac{r}{a} + k + 1, b\right)
 \end{aligned}$$

Where  $d(n, r)$  can be deduced from the relation

$$s(n, r) = (-1)^{n-r} d(n, r)$$

#### 2.10 Negative Moments

The negative moments of the LNK distribution is defined as:

$$E(X^{-r}) = \int_0^1 x^{-r} f(x; a, b) dx$$

By applying the substitution (4), we have

$$E(X^{-r}) = bc^{\frac{r}{a}} \sum_{k=0}^{\infty} \binom{n}{k} (-1)^k \left(1 - \frac{1}{c}\right)^k B\left(-\frac{r}{a} + k + 1, b\right)$$

**2.11 Incomplete Moments**

The Incomplete moments of the LNK distribution is defined as:

$$\begin{aligned} \phi_r(x) &= \int_0^x x^r f(x) dx \\ &= \int_0^x \frac{x^r abc x^{a-1} (1-x^a)^{b-1}}{[1-(1-c)x^a]^{1+b}} dx \\ &= \frac{b}{c^a} \sum_{k=0}^{\infty} \frac{\Gamma\left(\frac{r}{a} + k\right)}{\Gamma\left(\frac{r}{a}\right) k!} \left(1 - \frac{1}{c}\right)^k I\left(\frac{cx^a}{1-x^a + cx^a}; \frac{r}{a} + k + 1, b\right) \end{aligned}$$

**2.12 Scaled Total Time for Aging Properties**

The scaled total time of the LNK distribution is defined as:

$$\begin{aligned} S_F(F(x)) &= \frac{1}{\mu} \int_0^x S(y) dy = \frac{1}{\mu} \int_0^x \int_y^1 f(t) dt dy \\ S_F(F(x)) &= \frac{1}{\mu} \left( x - \int_0^x F(y) dy \right) \\ S_F(F(x)) &= \frac{1}{\mu} \int_0^x \left[ \frac{1-y^a}{1-(1-c)y^a} \right]^b dy \end{aligned}$$

**2.13 Conditional Moments**

The conditional moments of the LNK distribution is defined as:

$$\begin{aligned} M'_r &= E[X^r \mid X > x] \\ M'_r &= \frac{1}{\bar{F}_x} \int_x^1 t^r f(t) dt \end{aligned}$$

By applying the substitution (4) and after some simplification, we have

$$M'_r = \frac{(1-(1-c)x^a)^b}{(1-x^a)^b} \left[ \mu'_r - \frac{b}{c^a} \sum_{k=0}^{\infty} \frac{\Gamma\left(\frac{r}{a} + k\right)}{\Gamma\left(\frac{r}{a}\right) k!} \left(1 - \frac{1}{c}\right)^k I\left(\frac{cx^a}{1-x^a + cx^a}; \frac{r}{a} + k + 1, b\right) \right]$$

**2.14 Mean Residual Function**

The mean residual function of the LNK distribution is defined as:

$$m(x) = E[X - x | X > x]$$

$$m(x) = E[X | X > x] - x$$

$$= \frac{(1 - (1 - c)x^a)^b}{(1 - x^a)^b} \left[ \mu'_1 - \frac{b}{c^{\frac{1}{a}}} \sum_{k=0}^{\infty} \frac{\Gamma\left(\frac{1}{a} + k\right)}{\Gamma\left(\frac{1}{a}\right)k!} \left(1 - \frac{1}{c}\right)^k I\left(\frac{cx^a}{1 - x^a + cx^a}; \frac{1}{a} + k + 1, b\right) \right] - x$$

**2.15 Vitality Function**

The vitality function of the LNK distribution is defined as:

$$V(x) = E[X | X > x]$$

$$V(x) = \frac{(1 - (1 - c)x^a)^b}{(1 - x^a)^b} \left[ \mu'_1 - \frac{b}{c^{\frac{1}{a}}} \sum_{k=0}^{\infty} \frac{\Gamma\left(\frac{1}{a} + k\right)}{\Gamma\left(\frac{1}{a}\right)k!} \left(1 - \frac{1}{c}\right)^k I\left(\frac{cx^a}{1 - x^a + cx^a}; \frac{1}{a} + k + 1, b\right) \right]$$

**2.16 Geometric Vitality Function**

The Geometric vitality function of the LNK distribution is defined as:

$$\log G(x) = E[\log X | X > x]$$

$$\log G(x) = \frac{1}{1 - F(x)} \int_x^{\infty} \log t f(t) dt$$

$$= \frac{1}{1 - F(x)} \left( \int_0^x \log t f(t) dt - \int_0^x \log t f(t) dt \right)$$

$$= \frac{1}{1 - F(x)} \left( \left[ \frac{1}{a} (\psi(1) - \psi(1 + b)) - \frac{\ln c}{a} + b \sum_{i=1}^{\infty} \frac{1}{i} \left(1 - \frac{1}{c}\right)^i B(i + 1, b) \right] - \int_0^x \log t f(t) dt \right)$$

$$= \frac{1}{1 - F(x)} \left( \left[ \frac{1}{a} (\psi(1) - \psi(1 + b)) - \frac{\ln c}{a} + b \sum_{i=1}^{\infty} \frac{1}{i} \left(1 - \frac{1}{c}\right)^i B(i + 1, b) \right] - \frac{b}{a} \int_0^{\frac{cx^a}{1 - (1 - c)x^a}} (1 - u)^{b-1} (\ln u) du \right. \\ \left. + \frac{b}{a} \int_0^{\frac{cx^a}{1 - (1 - c)x^a}} (1 - u)^{b-1} \ln(c + u - uc) du \right)$$

### 2.17 Characteristics Function

The characteristic function of the LNK distribution is defined as:

$$\begin{aligned} \phi_x(t) &= E(e^{itx}) \\ &= \sum_{r=0}^{\infty} \frac{bc^{\frac{-r}{a}}}{\Gamma\left(\frac{r}{a}\right)} \sum_{k=0}^{\infty} \Gamma\left(\frac{r}{a}+k\right) \left(1-\frac{1}{c}\right)^k B\left(\frac{r}{a}+k+1, b\right) \frac{(it)^r}{k!r!} \quad \because e^x = \sum_{r=0}^{\infty} \frac{x^r}{r!} \end{aligned}$$

### 2.18 Information generating function

The information generating function of the LNK distribution is defined by  $P(s)$  and found from

$$\begin{aligned} P(s) &= E(f^{s-1}) \\ &= \int_0^1 \left( abc \frac{x^{a-1} (1-x^a)^{b-1}}{[1-(1-c)x^a]^{b+1}} \right)^s dx \end{aligned}$$

By applying the substitution (4), we have

$$\begin{aligned} P(s) &= a^{s-1} b^s c^{(s-1)/a} \sum_{i=0}^{\infty} \frac{\Gamma((a+1)(1-s)/a+i)}{\Gamma((a+1)(1-s)/a)i!} \left(1-\frac{1}{c}\right)^i \int_0^1 u^{(a-1)(s-1)/a} (1-u)^{s(b-1)} u^i du \\ &= a^{s-1} b^s c^{(s-1)/a} \sum_{i=0}^{\infty} \frac{\Gamma((a+1)(1-s)/a+i)}{\Gamma((a+1)(1-s)/a)i!} \left(1-\frac{1}{c}\right)^i \int_0^1 u^{(a-1)(s-1)/a+i} (1-u)^{s(b-1)} du \\ &= a^{s-1} b^s c^{(s-1)/a} \sum_{i=0}^{\infty} \frac{\Gamma((a+1)(1-s)/a+i)}{\Gamma((a+1)(1-s)/a)i!} \left(1-\frac{1}{c}\right)^i B((a-1)(s-1)/a+i+1, s(b-1)+1) \end{aligned}$$

### 2.19 Some Other Measures of Averages

#### 2.19.1 Harmonic Mean

This section contains Harmonic mean of  $x$  for the LNK distribution

The Harmonic mean of  $x$  is as under:

$$H_x = \frac{1}{E\left(\frac{1}{x}\right)}$$

Consider  $E\left(\frac{1}{x}\right) = \int_0^1 \frac{1}{x} f(x) dx$

$$= abc \int_0^1 \frac{x^{a-2} (1-x^a)^{b-1}}{[1-(1-c)x^a]^{1+b}} dx$$

By substituting  $x^a = y$ , we have

$$= \int_0^1 \frac{bcy^{-\frac{1}{a}} (1-y)^{b-1}}{[1-(1-c)y]^{1+b}} dy$$

After substitution (4), and simplification, we have

$$\begin{aligned} &= \int_0^1 \frac{bu^{-\frac{1}{a}} (1-u)^{b-1}}{[c+u-uc]^{-\frac{1}{a}}} du \\ &= \int_0^1 bu^{-\frac{1}{a}} (1-u)^{b-1} (c+u-uc)^{\frac{1}{a}} du \\ &= bc^{\frac{1}{a}} \int_0^1 u^{-\frac{1}{a}} (1-u)^{b-1} \sum_{i=0}^{\infty} \binom{1/a}{i} \left[ -\left(1-\frac{1}{c}\right)u \right]^i du \\ E\left(\frac{1}{x}\right) &= bc^{\frac{1}{a}} \sum_{i=0}^{\infty} \binom{1/a}{i} \left[ -\left(1-\frac{1}{c}\right) \right]^i B\left(\frac{-1}{a} + i + 1, b\right) \end{aligned}$$

$$H_x = \frac{1}{bc^{\frac{1}{a}} \sum_{i=0}^{\infty} \binom{1/a}{i} \left[ -\left(1-\frac{1}{c}\right) \right]^i B\left(\frac{-1}{a} + i + 1, b\right)} \quad \text{for } a > 1.$$

**Corollary**

If  $c=1$ , Then  $H_x = \frac{1}{bB\left(\frac{-1}{a} + 1, b\right)}$  is the Harmonic Mean of Kum distribution

### 2.19.2 Geometric Mean

The logarithm of the geometric  $G_X$  of a distribution with random variable  $X$  is the arithmetic mean of  $(\ln X)$ , or, equivalently, its expected value  $\ln G_X = E(\ln X)$ .

$$E(\ln x) = \int_0^1 (\ln x) f(x) dx$$



$$E(\ln x) = \int_0^1 (\ln x) \frac{abcx^{a-1}(1-x^a)^{b-1}}{[1-(1-c)x^a]^{1+b}} dx$$

By substitution  $x^a = y$  and (4), after simplification, we have

$$G_x = \exp \left[ \frac{1}{a} (\psi(1) - \psi(1+b)) - \frac{\ln c}{a} + b \sum_{i=1}^{\infty} \frac{1}{i} \left(1 - \frac{1}{c}\right)^i B(i+1, b) \right]$$

### 2.19.3 Mode

Mode is obtained by solving  $\frac{d}{dx} f(x) = 0$

$$f(x) = \frac{c^a x^{a-1} (1-x)^{b-1}}{B(a, b) [1-(1-c)x]^{a+b}}$$

Mode of LNK distribution is obtained by solving

$$\frac{d}{dx} f(x) = 0$$

$$f'(x) = \left\{ \frac{a-1}{x} - \frac{a(b-1)x^{a-1}}{(1-x^a)} - \frac{a(1+b)(c-1)x^{a-1}}{[1-(1-c)x^a]} \right\} f(x)$$

$$\frac{a-1}{x} - \frac{a(b-1)x^{a-1}}{(1-x^a)} - \frac{a(1+b)(c-1)x^{a-1}}{[1-(1-c)x^a]} = 0$$

Multiplying both sides by  $x$

$$a-1 - \frac{a(b-1)x^a}{(1-x^a)} - \frac{a(1+b)(c-1)x^a}{[1-(1-c)x^a]} = 0$$

Put  $x^a = y$

$$a-1 - \frac{a(b-1)y}{1-y} - \frac{a(1+b)(c-1)}{1-(1-c)y} = 0$$

$$(a-1)(1-y)(1-y+cy) - \{a(b-1)y(1-y+cy)\} - \{a(1+b)(c-1)y(1-y)\} = 0$$

$$-2a(1-c)y^2 + (2-c(ab+1))y + (a-1) = 0$$

$$Ay^2 + By + C = 0$$

$$y = \frac{-B \pm \sqrt{B^2 - 4AC}}{2A}$$

$$x = \left( \frac{-B \pm \sqrt{B^2 - 4AC}}{2A} \right)^{1/a}$$

Where  $y = x^a$

$$A = -2a(1-c)$$

$$B = 2 - c - abc$$

$$C = (a-1)$$

### 2.20 Points of Inflection

The points of the inflection of the LNK distribution are found from

$f''(x) = 0$  under  $f'''(x) \neq 0$  or equivalently as

$$\frac{d^2 \ln f(x)}{dx^2} = 0 \text{ and } \frac{d^3 \ln f(x)}{dx^3} \neq 0$$

Where as

$$\frac{d^2 \ln f(x)}{dx^2} = \frac{-(a-1)}{x^2} - \frac{a(b-1)x^{a-2}[(a-1) + x^a]}{(1-x^a)^2} + \frac{a(1+b)(1-c)x^{a-2}[(a-1) + (1-c)x^a]}{[1-(1-c)x^a]^2}$$

So

$$\frac{-(a-1)}{x^2} - \frac{a(b-1)x^{a-2}[(a-1) + x^a]}{(1-x^a)^2} + \frac{a(1+b)(1-c)x^{a-2}[(a-1) + (1-c)x^a]}{[1-(1-c)x^a]^2} = 0$$

After some Simplification

$$Ay^4 + By^3 + Cy^2 + Dy + E = 0$$

Where  $y = x^a$

$$A = -(a+1)(1-c)^2$$

$$B = (1-c)[(2-c)(a+2-a^2) - abc(1+a)]$$

$$C = (1-c)(4a^2 - 6) - abc(c-2) - c^2$$

$$D = (a-1)[(c-2)(a+2) + abc]$$

$$E = (a-1)$$

Table 2.1 Points of inflection of LNK distribution

A	B	C	Point (s) of inflection
1	2	0.1	0.8482
1	2	0.2	0.6585
1	2	0.3	0.4146
1	2	0.4	0.0894
1	3	0.1	0.7318
1	3	0.2	0.3965
2	3	0.1	0.4432 , 0.8678
2	3	0.2	0.5681 , 0.6607
2	4	0.1	0.4625, 0.8089
2	5	0.1	0.4888, 0.7454
2	6	0.1	0.5342, 0.6644
3	4	0.1	0.5492, 0.8741
3	5	0.1	0.5665, 0.8335
3	6	0.1	0.5897, 0.7881
3	7	0.1	0.6294, 0.7273

**2.21 Bonferroni Curve**

The Bonferroni curve of the LNK distribution is as:

$$B_F(F(x)) = \frac{1}{\mu F(x)} \int_0^x y f(y) dy$$

$$B_F(F(x)) = \frac{\sum_{k=0}^{\infty} \frac{\Gamma\left(\frac{1}{a} + k\right)}{\Gamma\left(\frac{1}{a}\right) k!} \left(1 - \frac{1}{c}\right)^k I\left(\frac{cx^a}{1-x^a + cx^a}; \frac{1}{a} + k + 1, b\right)}{\left(1 - \left(\frac{1-x^a}{1-(1-c)x^a}\right)^b\right) \sum_{k=0}^{\infty} \frac{\Gamma\left(\frac{1}{a} + k\right)}{\Gamma\left(\frac{1}{a}\right) k!} \left(1 - \frac{1}{c}\right)^k B\left(\frac{1}{a} + k + 1, b\right)}$$

**2.22 Lorenz Curve**

The Lorenz curve of the LNK distribution is as:

$$L(F(x)) = F(x)B(F(x))$$

$$L(F(x)) = \frac{\sum_{k=0}^{\infty} \frac{\Gamma\left(\frac{1}{a} + k\right)}{\Gamma\left(\frac{1}{a}\right) k!} \left(1 - \frac{1}{c}\right)^k I\left(\frac{cx^a}{1-x^a + cx^a}; \frac{1}{a} + k + 1, b\right)}{\sum_{k=0}^{\infty} \frac{\Gamma\left(\frac{1}{a} + k\right)}{\Gamma\left(\frac{1}{a}\right) k!} \left(1 - \frac{1}{c}\right)^k B\left(\frac{1}{a} + k + 1, b\right)}$$

**2.23 Gini Coefficient**

The Gini coefficient of the LNK distribution is explained as:

$$G = 1 - \mu^{-1} \int_0^1 (1 - F(x))^2 dx$$

$$= 1 - \frac{1}{\mu} \int_0^1 \left(\frac{1 - x^a}{1 - (1 - c)x^a}\right)^{2b} dx$$

**2.24 Mean deviations**

The mean deviation of the LNK distribution from arithmetic mean and median are denoted by  $\delta_1(Y)$  and  $\delta_2(Y)$  respectively and are found from

$$\delta_1(Y) = \int_0^1 |y - \mu| f(y) dy$$

$$= \int_0^{\mu} \mu f(y) dy - \int_0^{\mu} y f(y) dy + \int_{\mu}^1 y f(y) dy - \int_{\mu}^1 \mu f(y) dy$$

$$= 2\mu F(\mu) - 2J(\mu)$$

$$= 2\mu \left(1 - \left(\frac{1 - \mu^a}{1 - (1 - c)\mu^a}\right)^b\right) - 2 \frac{b}{c^a} \sum_{k=0}^{\infty} \frac{\Gamma\left(\frac{1}{a} + k\right)}{\Gamma\left(\frac{1}{a}\right) k!} \left(1 - \frac{1}{c}\right)^k I\left(\frac{c\mu^a}{1 - \mu^a + c\mu^a}; \frac{1}{a} + k + 1, b\right)$$

The mean deviation about median

$$\delta_2(Y) = \int_0^1 |y - M| f(y) dy$$

$$= \int_0^M M f(y) dy - \int_0^M y f(y) dy + \int_M^1 y f(y) dy - \int_M^1 M f(y) dy$$

$$\begin{aligned}
&= \mu - 2J(M) \\
&= \mu - 2 \frac{b}{c^a} \sum_{k=0}^{\infty} \frac{\Gamma\left(\frac{1}{a} + k\right)}{\Gamma\left(\frac{1}{a}\right) k!} \left(1 - \frac{1}{c}\right)^k I\left(\frac{cM^a}{1 - M^a + cM^a}; \frac{1}{a} + k + 1, b\right)
\end{aligned}$$

### 3. ESTIMATION

Here, we consider

#### 3.1 Maximum-likelihood estimation

The likelihood function of random sample  $x_1, x_2, \dots, x_n$  of observation is given by

$$f(x_1, x_2, \dots, x_n; a, b, c) = \prod_{i=1}^n f(x_i, a, b, c)$$

$$L(x_i, a, b, c) = \prod_{i=1}^n f(x_i, a, b, c)$$

$$L(x_i, a, b, c) = \frac{(abc)^n \left(\prod_{i=1}^n x_i\right)^{a-1} \left(\prod_{i=1}^n (1 - x_i^a)\right)^{b-1}}{\left(\prod_{i=1}^n (1 - (1-c)x_i^a)\right)^{1+b}}$$

Taking ln both sides

$$\ln L = n \ln abc + (a-1) \sum_{i=1}^n \ln x_i + (b-1) \sum_{i=1}^n \ln(1 - x_i^a) - (1+b) \sum_{i=1}^n \ln(1 - (1-c)x_i^a)$$

It follows that the maximum-likelihood estimates  $(a, b, c)$ , say  $(\hat{a}, \hat{b}, \hat{c})$ , are the simultaneous solutions of the equations:

$$\frac{\partial L}{\partial a} = \frac{n}{a} + \sum_{i=1}^n \ln x_i - (b-1) \sum_{i=1}^n \frac{x_i^a \ln x_i}{(1 - x_i^a)} + (1+b) \sum_{i=1}^n \frac{(1-c)x_i^a \ln x_i}{(1 - (1-c)x_i^a)}$$

$$\frac{\partial L}{\partial b} = \frac{n}{b} + \sum_{i=1}^n \ln(1 - x_i^a) - \sum_{i=1}^n \ln(1 - (1-c)x_i^a)$$

$$\frac{\partial L}{\partial c} = \frac{n}{c} - (1+b) \sum_{i=1}^n \frac{x_i^a}{1 - (1-c)x_i^a}$$

For the ML estimators, we solve the following three equations simultaneously we have

$$\frac{\partial L}{\partial a} = 0, \frac{\partial L}{\partial b} = 0, \frac{\partial L}{\partial c} = 0$$

Since the equations are of implicit and have the complicated forms. Therefore, for numerically solution of the equations we can use different softwares like The R Mathematica 10.2 etc.

#### 4. SIMULATION STUDY

**Table 1** The Bias, MSE values for the LNK model when a=b=c=2.0

	<b>n=10</b>	<b>n=20</b>	<b>n=30</b>	<b>n=50</b>	<b>n=70</b>	<b>n=100</b>
Bias(a)	0.6789	0.2966	0.1591	0.0732	0.0398	0.0134
Bias(b)	148.4399	98.1754	80.9118	56.4929	42.0678	37.0448
Bias(c)	519.4312	63.2427	5.1913	0.3873	-0.1171	-0.6091
MSE(a)	2.3026	0.7537	0.3908	0.1723	0.1132	0.0843
MSE(b)	45131.8214	22559.3813	17215.8524	9705.1161	5973.0134	5221.2081
MSE(c)	4729852.1234	592121.2123	17854.2713	33.3845	15.21264	4.6655
	<b>n=150</b>	<b>n=200</b>	<b>n=300</b>	<b>n=500</b>	<b>n=700</b>	<b>n=1000</b>
Bias(a)	0.0165	0.0013	0.0025	-0.0071	-0.0157	-0.0134
Bias(b)	25.9659	17.9847	12.0603	6.9705	5.6428	4.7767
Bias(c)	-0.7730	-0.9022	-0.9606	-0.9643	-0.9604	-0.9707
MSE(a)	0.0578	0.0445	0.0286	0.0179	0.0123	0.0081
MSE(b)	2688.8471	1429.0832	685.7248	176.8713	80.2722	36.4003
MSE(c)	2.2229	1.8662	1.5139	1.3256	1.2007	1.1071

**Table 2** The Bias, MSE values for the LNK model when  $a=b=c=2.5$ 

	<b>n=10</b>	<b>n=20</b>	<b>n=30</b>	<b>n=50</b>	<b>n=70</b>	<b>n=100</b>
Bias(a)	0.7101	0.1423	0.0226	-0.0336	-0.0764	-0.0892
Bias(b)	148.6944	83.5053	70.3768	49.9135	40.9383	28.8218
Bias(c)	908.6814	240.0717	101.7437	6.0567	3.3570	1.6298
MSE(a)	2.5889	0.6457	0.4488	0.2476	0.1627	0.1103
MSE(b)	46025.2324	19184.2612	13459.1232	7983.5764	5993.6893	3566.5871
MSE(c)	8932954	2172753	858508	1049.3921	181.7757	45.5042
	<b>n=150</b>	<b>n=200</b>	<b>n=300</b>	<b>n=500</b>	<b>n=700</b>	<b>n=1000</b>
Bias(a)	-0.1165	-0.1040	-0.1170	-0.1267	-0.1286	-0.1309
Bias(b)	16.4083	11.5937	8.2968	4.7545	3.6887	3.1612
Bias(c)	0.9505	1.0242	0.6185	0.4136	0.4096	0.2923
MSE(a)	0.0804	0.0619	0.0491	0.0373	0.0310	0.0282
MSE(b)	1369.2381	791.9238	402.1093	108.3226	41.8772	16.9204
MSE(c)	15.8173	13.1791	6.4383	3.0282	2.2463	1.4012

**Table 3** The Bias, MSE values for the LNK model when  $a=b=c=3.0$ 

	<b>n=10</b>	<b>n=20</b>	<b>n=30</b>	<b>n=50</b>	<b>n=70</b>	<b>n=100</b>
Bias(a)	0.4603	0.0535	-0.0688	-0.1884	-0.2438	-0.2725
Bias(b)	139.5197	81.7559	59.6459	39.1355	28.2413	16.5283
Bias(c)	1318.3864	540.5148	167.8884	61.5224	12.9443	9.0277
MSE(a)	2.1234	0.9155	0.6307	0.3236	0.2394	0.2077
MSE(b)	43208.6434	18732.8834	11124.5823	5996.8751	3686.9514	2041.9723
MSE(c)	12800727	4713924	1058158	285476.4	1760.2651	455.9306
	<b>n=150</b>	<b>n=200</b>	<b>n=300</b>	<b>n=500</b>	<b>n=700</b>	<b>n=1000</b>
Bias(a)	-0.2895	-0.3107	-0.3037	-0.3063	-0.3198	-0.3133
Bias(b)	10.2456	6.4328	3.8052	2.2502	1.5573	1.3494
Bias(c)	7.5233	7.4066	5.9516	5.0419	5.2533	4.9514
MSE(a)	0.1698	0.1596	0.1384	0.1194	0.1204	0.1101
MSE(b)	845.9625	368.9467	139.7221	35.0732	8.4393	3.1515
MSE(c)	238.0884	154.0213	83.8912	46.6563	42.8393	34.4719

### 5. SOME TRANSFORMED DISTRIBUTIONS

5.1 Following table show the different transformation of LNK distribution:

Sr. No.	a	b	c	Transformation	Resulting Distribution
1	a	b	1	X	Kumaraswamy distribution (1980)
2	a	b	1	1 - x	Corderio (2012) distribution $F(x) = 1 - \left(1 - (1-x)^a\right)^b$ ; $0 < x < 1, a, b > 0$
3	1	b	1	X	B(1,b) and Kum(1,b) Power distribution
4	a	1	1	X	B(a,1) and Kum (a,1) Power distribution
5	a	1	1	-logx	Exp(a)
6	1	b	1	-log(1-x)	Exp(b)
7	a	b	1	X	Kum(a,b) or $G B_1(a,1,1,b)$
8	1	b	1	$\frac{x}{1+x}$	$F(x) = 1 - \frac{1}{(1+x)^b}$ $0 < x < \infty, b > 0$
9	1	b	c	$\frac{cx}{1-x+cx}$	$F(x) = 1 - (1-x)^b$ ; $0 < x < 1, b > 0$
10	1	1	1	X	Uniform (0,1)

### 5.2 Asymptotes and Shapes

The asymptotes of (1.1), (1.2) and (3.1) as  $x \rightarrow 0, 1$  are given by

	pdf $f(x)$	hrf $h(x)$
$x \rightarrow 0$	$abc x^{a-1}$ where $a, b, c < 1$	$abc x^{a-1}$ where $a, b, c < 1$
$x \rightarrow 1$	$abc (1-x^a)^{b-1}$ where $a, b, c < 1$	$\frac{abc}{(1-x^a)}$ where $a, b, c < 1$



## 6. APPLICATION

In order to prove that LNK distribution can be a better model than the Power distribution, Beta distribution with ( $a = 1$ ), Beta distribution, Kumaraswamy distribution, let us use three real data sets.

The following tables show the numerical values with MLEs and their corresponding standard errors (in parentheses) of the model parameters including loglikelihood, Kolmogorov-Smirnov test (KS), Akaike information criterion (AIC) and Consistent Akaike information criterion (CAIC) for comparing LNK distribution with the Power distribution, Beta distribution with ( $a = 1$ ), Beta distribution, Kumaraswamy distribution. It is quite evident from the reports that LNK distribution is better. The plots of the fitted distributions to real datasets are shown in figures.

### Data Set 1:

The following right to skewed dataset presented by Cordeiro and Brito (2012) [13] is obtained from the measurements on petroleum rock samples. The data consists of 48 rock samples from a petroleum reservoir. The dataset corresponds to twelve core samples from petroleum reservoirs that were sampled by four cross-sections. Each core sample was measured for permeability and each cross-section has the following variables: the total area of pores, the total perimeter of pores and shape. We analyze the shape perimeter by squared (area) variable and the observations are:

0.0903296,	0.2036540,	0.2043140,	0.2808870,	0.1976530,	0.3286410,
0.1486220,	0.1623940,	0.2627270,	0.1794550,	0.3266350,	0.2300810,
0.1833120,	0.1509440,	0.2000710,	0.1918020,	0.1541920,	0.4641250,
0.1170630,	0.1481410,	0.1448100,	0.1330830,	0.2760160,	0.4204770,
0.1224170,	0.2285950,	0.1138520,	0.2252140,	0.1769690,	0.2007440,
0.1670450,	0.2316230,	0.2910290,	0.3412730,	0.4387120,	0.2626510,
0.1896510,	0.1725670,	0.2400770,	0.3116460,	0.1635860,	0.1824530,
0.1641270,	0.1534810,	0.1618650,	0.2760160,	0.2538320,	0.2004470.

### Data set 2

The symmetric behavior of the following dataset, discussed by Dasgupta (2011) [14], consists of 50 observations relates to holes operation on jobs made of iron sheet. This dataset is as follows:

0.04, 0.02, 0.06, 0.12, 0.14, 0.08, 0.22, 0.12, 0.08, 0.26, 0.24, 0.04, 0.14, 0.16, 0.08, 0.26, 0.32, 0.28, 0.14, 0.16, 0.24, 0.22, 0.12, 0.18, 0.24, 0.32, 0.16, 0.14, 0.08, 0.16, 0.24, 0.16, 0.32, 0.18, 0.24, 0.22, 0.16, 0.12, 0.24, 0.06, 0.02, 0.18, 0.22, 0.14, 0.06, 0.04, 0.14, 0.26, 0.18, 0.16.

**Data Set 3:**

The following second data set presented by Cordeiro and Brito (2012) [13], displays the skewed symmetric trend of data. This data discusses the total milk production in the first birth of 107 cows from SINDI race. These cows are property of the Carnaúba farm which belongs to the *Agropecuária Manoel Dantas Ltda* (AMDA), located in Taperoá City, Paraíba (Brazil). The data of proportion of total milk is as under:

0.4365, 0.4260, 0.5140, 0.6907, 0.7471, 0.2605, 0.6196, 0.8781, 0.4990, 0.6058, 0.6891, 0.5770, 0.5394, 0.1479, 0.2356, 0.6012, 0.1525, 0.5483, 0.6927, 0.7261, 0.3323, 0.0671, 0.2361, 0.4800, 0.5707, 0.7131, 0.5853, 0.6768, 0.5350, 0.4151, 0.6789, 0.4576, 0.3259, 0.2303, 0.7687, 0.4371, 0.3383, 0.6114, 0.3480, 0.4564, 0.7804, 0.3406, 0.4823, 0.5912, 0.5744, 0.5481, 0.1131, 0.7290, 0.0168, 0.5529, 0.4530, 0.3891, 0.4752, 0.3134, 0.3175, 0.1167, 0.6750, 0.5113, 0.5447, 0.4143, 0.5627, 0.5150, 0.0776, 0.3945, 0.4553, 0.4470, 0.5285, 0.5232, 0.6465, 0.0650, 0.8492, 0.8147, 0.3627, 0.3906, 0.4438, 0.4612, 0.3188, 0.2160, 0.6707, 0.6220, 0.5629, 0.4675, 0.6844, 0.3413, 0.4332, 0.0854, 0.3821, 0.4694, 0.3635, 0.4111, 0.5349, 0.3751, 0.1546, 0.4517, 0.2681, 0.4049, 0.5553, 0.5878, 0.4741, 0.3598, 0.7629, 0.5941, 0.6174, 0.6860, 0.0609, 0.6488, 0.2747.

**Estimated Parameters by MLE with their S.E. and Goodness of Fit****Data set 1**

Model	$\hat{\alpha}$	$\hat{\beta}$	$\hat{\gamma}$	$-\ln(L)$	KS	AIC	CAIC
Power	0.63 (0.0909)			-6.0118	0.4295	14.0237	14.1107
Beta ( $\alpha = 1$ )		3.9643 (0.5722)		- 30.2205	0.9156	62.4412	62.5281
Kumaraswamy	44.6597 (17.5894)	2.7187 (0.2937)		- 52.4915	0.1533	108.9831	109.2497
Beta	5.9417 (1.1825)	21.2057 (4.3513)		- 55.6002	0.1427	115.2004	115.4671
Libby Novick Kumaraswamy	5.5040 (0.6078)	0.7469 (0.2042)	10.4007 (130.7012)	- 57.7939	0.0852	121.588	122.1334

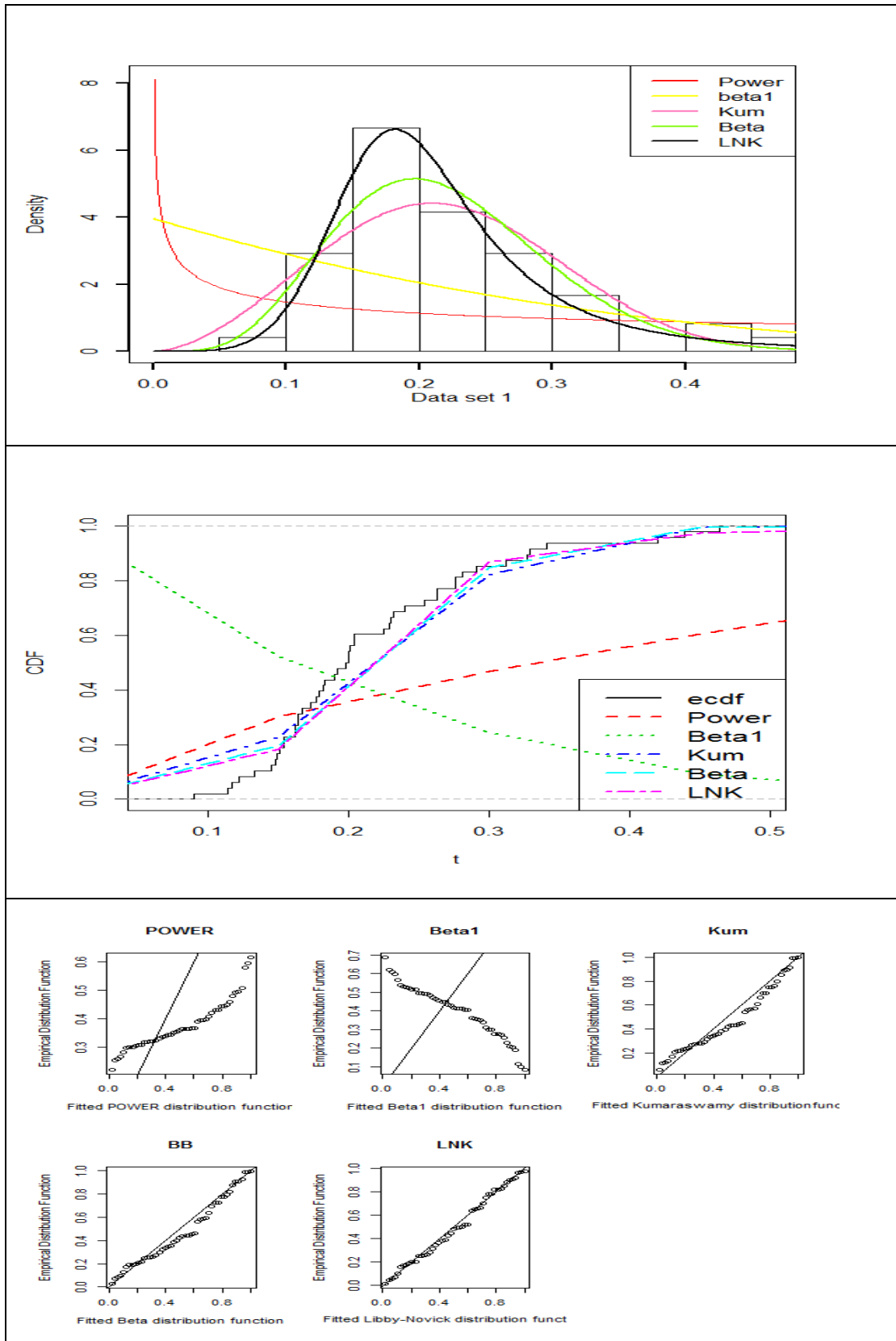
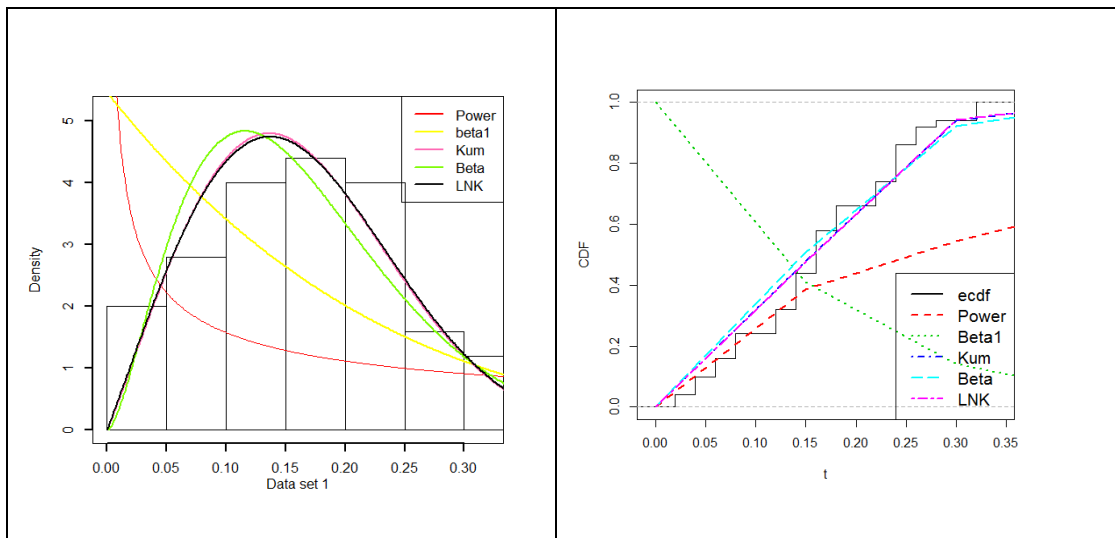


Figure 6.1: pdf, cdf and Q-Q graphs of the densities for data set 1

**Estimated Parameters by MLE with their S.E. and Goodness of Fit**

**Data set 2**

Model	$\hat{\alpha}$	$\hat{\beta}$	$\hat{\gamma}$	$-\ln(L)$	KS	AIC	CAIC
Power	0.5033 (0.0711)			-15.0139	0.4364	32.0279	32.1113
Beta ( $\alpha = 1$ )		5.4693 (0.7735)		-44.0998	0.8786	90.1997	90.2831
Kumaraswamy	33.1367 (13.9342)	2.0773 (0.2551)		-56.0686	0.0902	116.1374	116.3927
Beta	2.6825 (0.5074)	13.8656 (2.8295)		-54.6066	0.1214	113.2133	113.4686
Libby Novick Kumaraswamy	2.0376 (0.2687)	226.0664 (980.0583)	0.1342 (0.5945)	-56.2244	0.0900	118.4489	118.9707



**Figure 6.2:** pdf and cdf graphs of the competitors distributions for data set 2

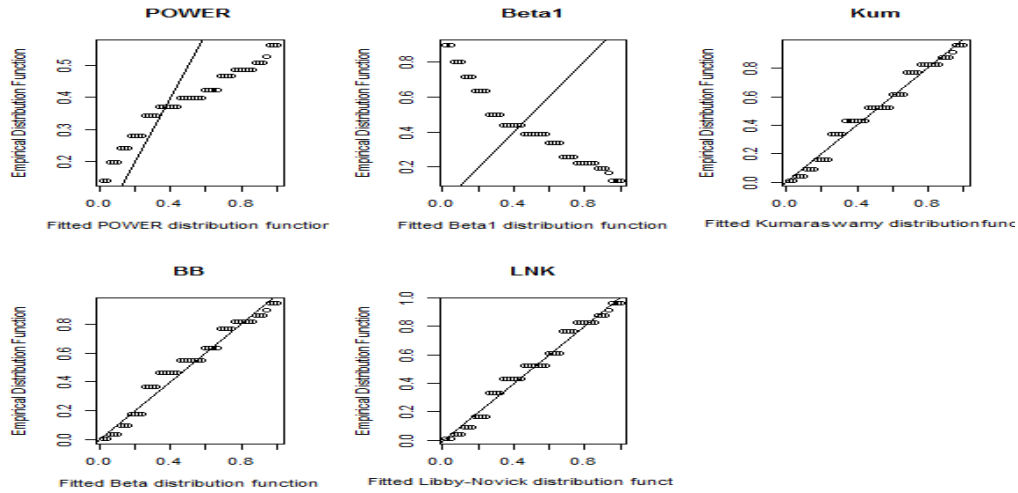


Figure 6.3: Q-Q plots of the densities for data set 2

Estimated Parameters by MLE with their S.E. and Goodness of Fit data3

Model	$\hat{\alpha}$	$\hat{\beta}$	$\hat{\gamma}$	$-\ln(L)$	KS	AIC	CAIC
Power	1.1123 (0.1075)			-0.58545	0.2418	3.17090	3.209
Beta ( $\alpha = 1$ )		1.4181 (0.1371)		-5.8304	0.9669	13.6608	13.6989
Kumaraswamy	3.4363 (0.5822)	2.1948 (0.2224)		-25.3946	0.0769	54.7893	54.9047
Beta	2.4125 (0.3145)	2.8296 (0.3745)		-23.7772	0.0816	51.5544	51.6698
Libby Novick Kumaraswamy	1.7588 (1.0411)	18.3624 (201.0970)	0.1044 (1.2976)	-27.2446	0.0698	60.4892	60.7222

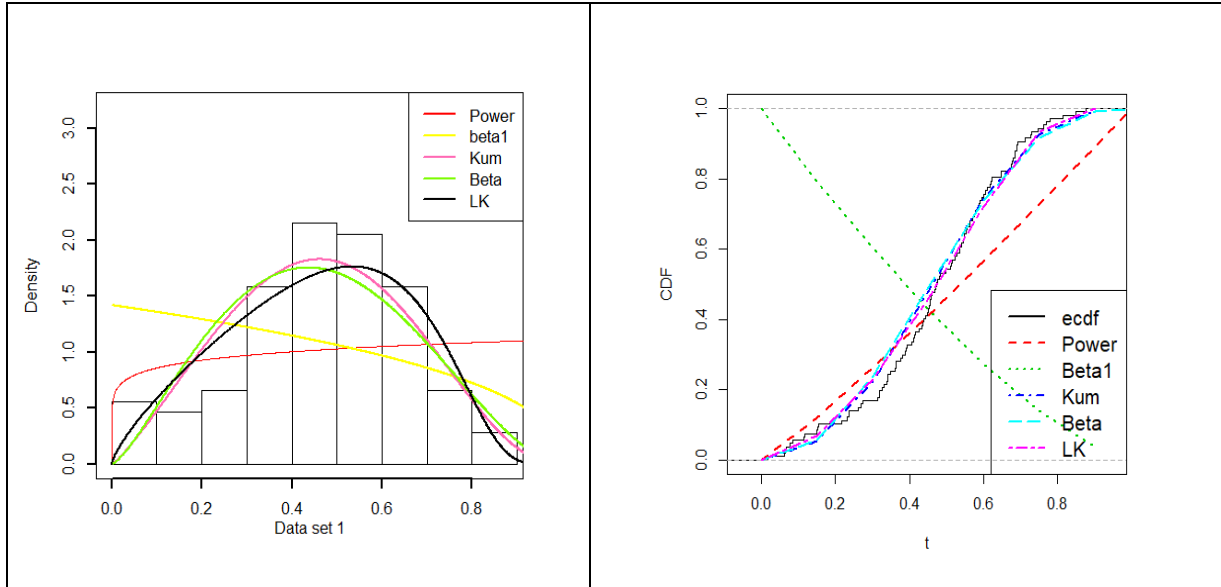


Figure 6.4: pdf and cdf graphs of the competitors distributions for data set 3

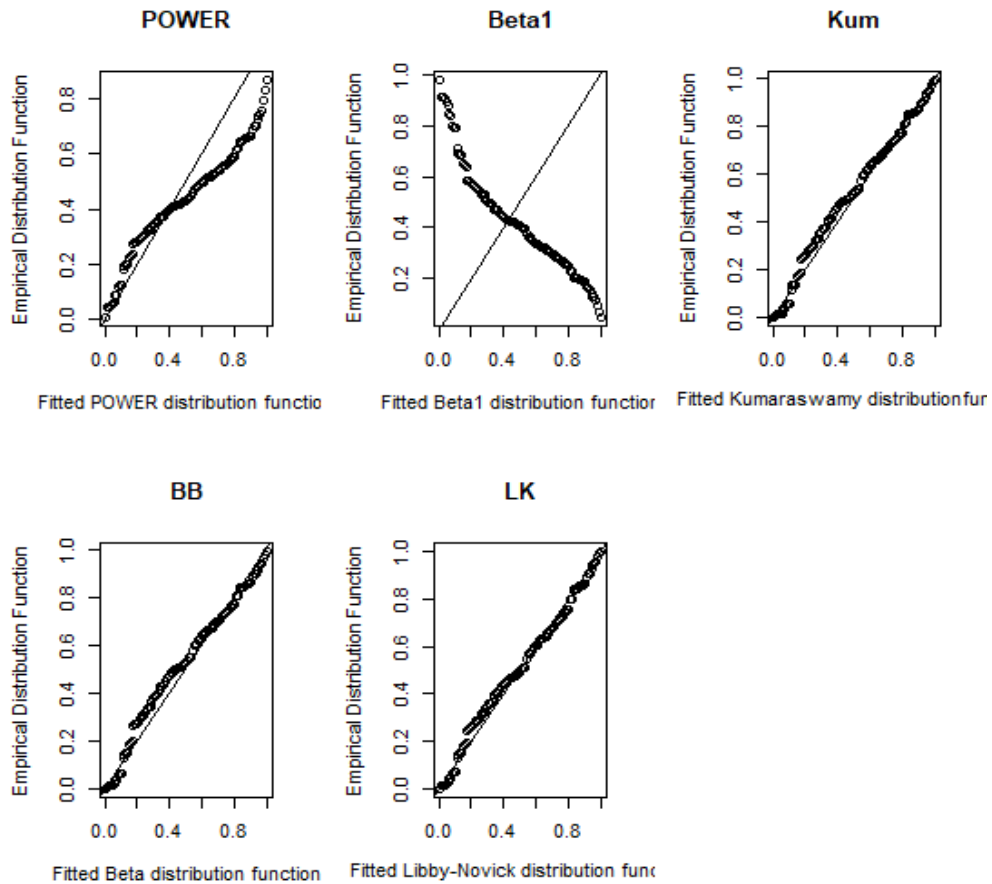


Figure 6.5: Q-Q plots of the densities for data set 3

## 6 CONCLUSION

The LNK distribution has some properties like the LNB distribution but, there are numerous benefits of the LNK distribution over LNB distribution: the LNK distribution has simple closed form of both its *cdf* and quantile function and that's why it is easy to use in simulation studies. The distribution and quantile function of LNK do not involve any special functions. To compare the proposed model with other models, we apply these models to three sets of real data from different fields of science and engineering and it is examined by using well-known statistics. We conclude that the LNK distribution is better than the power, Beta with ( $a=1$ ), Beta distribution and Kumaraswamy distribution.

**Appendix 1: Quantile values of LNK distribution**

a	b	c	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9	0.99	0.999
0.5	0.5	0.1	0.49155	0.72089	0.83237	0.89632	0.93652	0.96296	0.98050	0.99171	0.99798	0.99998	1
0.5	0.5	0.2	0.29135	0.54420	0.70361	0.80797	0.87890	0.92795	0.96158	0.98353	0.99597	0.99996	1
0.5	0.5	0.6	0.07899	0.23413	0.40237	0.55900	0.69444	0.80539	0.89110	0.95181	0.98798	0.99988	0.99999
0.5	0.5	1.0	0.03610	0.12960	0.26010	0.40960	0.56250	0.70560	0.82810	0.92160	0.98010	0.99980	0.99999
0.5	0.5	1.5	0.01828	0.07438	0.16780	0.29416	0.44444	0.60493	0.75831	0.88581	0.97037	0.99970	0.99999
0.5	0.5	3.0	0.00525	0.02493	0.06634	0.13845	0.25000	0.40495	0.59472	0.79012	0.94204	0.99940	0.99999
0.5	0.5	6.0	0.00141	0.00734	0.02185	0.05224	0.11111	0.21777	0.39386	0.64000	0.88898	0.99880	0.99998
0.5	0.5	10	0.00052	0.00283	0.00888	0.02278	0.05325	0.11851	0.25277	0.49827	0.82493	0.99800	0.99998
1.0	1.0	0.1	0.52631	0.71428	0.81081	0.86956	0.90909	0.93750	0.95890	0.97561	0.98901	0.99899	0.99990
1.0	1.0	0.2	0.35714	0.55555	0.68181	0.76923	0.83333	0.88235	0.92105	0.95238	0.97826	0.99798	0.99980
1.0	1.0	0.6	0.15625	0.29411	0.41666	0.52631	0.62500	0.71428	0.79545	0.86956	0.93750	0.99397	0.99940
1.0	1.0	1.0	0.10000	0.20000	0.30000	0.40000	0.50000	0.60000	0.70000	0.80000	0.90000	0.99000	0.99900
1.0	1.0	1.5	0.06896	0.14285	0.22222	0.30769	0.40000	0.50000	0.60869	0.72727	0.85714	0.98507	0.99850
1.0	1.0	3.0	0.03571	0.07692	0.12500	0.18181	0.25000	0.33333	0.43750	0.57142	0.75000	0.97058	0.99700
1.0	1.0	6.0	0.01818	0.04000	0.06666	0.10000	0.14285	0.20000	0.28000	0.40000	0.60000	0.94285	0.99403
1.0	1.0	10.0	0.01098	0.02439	0.04109	0.06250	0.0909	0.13043	0.18918	0.28571	0.47368	0.90825	0.99008
2.0	2.0	0.1	0.59248	0.73576	0.81319	0.86269	0.89751	0.92368	0.94444	0.96184	0.97764	0.99449	0.998371
2.0	2.0	0.2	0.46139	0.60921	0.70282	0.76984	0.82120	0.86253	0.89722	0.92775	0.95673	0.98907	0.99675
2.0	2.0	0.6	0.28757	0.40544	0.49548	0.57148	0.63906	0.70143	0.76103	0.82049	0.88475	0.96824	0.99034



2.0	2.0	1.0	0.22653	0.32492	0.40415	0.47476	0.54119	0.60625	0.67251	0.74349	0.82690	0.94868	0.98406
2.0	2.0	1.5	0.18656	0.27009	0.33935	0.40308	0.46517	0.52843	0.59585	0.67213	0.76838	0.92582	0.97637
2.0	2.0	3.0	0.13308	0.19456	0.24718	0.29735	0.34831	0.40283	0.46458	0.54018	0.64719	0.86602	0.95434
2.0	2.0	6.0	0.09452	0.13889	0.17751	0.21507	0.25412	0.29715	0.34781	0.41330	0.51469	0.77459	0.91442
2.0	2.0	10.0	0.07334	0.10800	0.13838	0.16815	0.19943	0.23435	0.27618	0.33167	0.42164	0.68824	0.86823
2.5	2.5	1.0	0.27941	0.37373	0.44615	0.50896	0.56705	0.62340	0.68063	0.74227	0.81622	0.93330	0.97426
3.0	3.0	0.1	0.64095	0.75811	0.82327	0.86619	0.89717	0.9210	0.94038	0.95702	0.97268	0.99101	0.99632
3.0	3.0	0.2	0.53323	0.65307	0.72871	0.78371	0.82677	0.86225	0.89283	0.92060	0.94813	0.98233	0.99270
3.0	3.0	0.6	0.38309	0.48491	0.55810	0.61822	0.67111	0.71995	0.76713	0.81531	0.86978	0.95043	0.97871
3.0	3.0	1.0	0.32557	0.41540	0.48216	0.53897	0.59088	0.64085	0.69143	0.74602	0.81222	0.92230	0.965489
3.0	3.0	1.5	0.28551	0.36582	0.42659	0.47932	0.52858	0.57723	0.62799	0.68488	0.75764	0.89138	0.94991
3.0	3.0	3.0	0.22749	0.29276	0.34308	0.38769	0.43040	0.47385	0.52089	0.57626	0.65255	0.81847	0.90856
3.0	3.0	6.0	0.18092	0.23334	0.27416	0.31075	0.34627	0.38301	0.42366	0.47297	0.54441	0.72284	0.84343
3.0	3.0	10.0	0.15271	0.19714	0.23187	0.26316	0.29369	0.32550	0.36102	0.40471	0.46950	0.64388	0.77952

### Appendix 2: Variance, Skewness and Kurtosis of LNK distribution

Following table shows the skewness and kurtosis alongwith variance of LNK distribution

a	b	c	Variance	Skewness	Kurtosis
0.5	0.5	0.1	0.05296	-1.90608	6.01753
0.5	0.5	0.2	0.07621	-1.27182	3.53694
0.5	0.5	0.6	0.11146	-0.46476	1.80245
0.5	0.5	1	0.12190	-0.13530	1.5386
0.5	0.5	1.5	0.12589	0.11638	1.50603
0.5	0.5	3.0	0.12346	0.54039	1.77633
0.5	0.5	6.0	0.11145	0.97311	2.46619
0.5	0.5	10.0	0.09890	1.30763	3.28243
1.0	1.0	0.1	0.04264	-1.82849	5.97880
1.0	1.0	0.2	0.05954	-1.19142	3.56560
1.0	1.0	0.6	0.08049	-0.35626	1.95722
1.0	1.0	1	0.08333	0.00000	1.80000
1.0	1.0	1.5	0.08152	0.28209	1.89855
1.0	1.0	3.0	0.07109	0.78491	2.56442
1.0	1.0	6.0	0.05508	1.34776	4.06188
1.0	1.0	10.0	0.04264	1.82849	5.97880
1.5	1.5	0.1	0.03501	-87.1461	386.365
1.5	1.5	0.2	0.04780	-40.6365	139.708
1.5	1.5	0.6	0.06162	-13.6714	32.6900
1.5	1.5	1	0.06205	-0.04546	2.00878
1.5	1.5	1.5	0.05922	0.21997	2.1023
1.5	1.5	3.0	0.04919	0.68276	2.73117
1.5	1.5	6.0	0.03603	1.18148	4.11004
1.5	1.5	10.0	0.02662	1.58988	5.82322
2.0	2.0	0.1	0.02935	-1.80078	6.22199
2.0	2.0	0.2	0.03939	-1.20368	3.91541
2.0	2.0	0.6	0.04923	-0.43784	2.35643
2.0	2.0	1	0.04888	-0.12530	2.18005
2.0	2.0	3.0	0.03789	0.50497	2.69308
2.0	2.0	6.0	0.02764	0.89112	3.64976
2.0	2.0	10.0	0.02052	1.17422	4.71545
2.5	2.5	0.1	0.02502	-694.098	3782.96
2.5	2.5	0.2	0.03311	-273.261	1177.61
2.5	2.5	0.6	0.04050	-278.916	868.895
2.5	2.5	1.5	0.03761	0.00355	2.3443
2.5	2.5	3.0	0.03094	0.33237	2.64818
2.5	2.5	6.0	0.02295	0.62511	3.25548
2.5	2.5	10.0	0.01744	0.81623	3.85719

**Conflicts of Interest:** The author declares that there are no conflicts of interest regarding the publication of this paper.

## References

- [1] Z. Iqbal, A. Saboor, M. Ahmad, A note on Libby-Novick Kumaraswamy distribution. Presented in 15<sup>th</sup> ISSOS conference (2017).
- [2] D. L. Libby, M. R. Novick, Multivariate generalized beta distributions with applications to utility assessment. *J. Educ. Stat.* 7(4) (1982), 271-294.
- [3] P. Kumaraswamy, A generalized probability density function for double-bounded random processes. *J. Hydrol.* 46(1-2) (1980), 79-88.
- [4] M. C. Jones, Kumaraswamy's distribution: A beta-type distribution with some tractability advantages. *Stat. Methodol.* 6(1) (2009), 70-81.
- [5] J. McDonald, Some generalized functions for the size distribution of income. *Econometrica.* 52(3) (1984), 647-665.
- [6] P. A. Mitnik, New properties of the Kumaraswamy distribution. *Commun. Stat., Theory Meth.* 42(5) (2013), 741-755.
- [7] J. J. Chen, M. R. Novick, Bayesian analysis for binomial models with generalized beta prior distributions. *J. Educ. Stat.* 9(2) (1984), 163-175.
- [8] M. M. Ristić, B. V. Popović, S. Nadarajah, Libby and Novick's generalized beta exponential distribution. *J. Stat. Comput. Simul.* 85(4) (2013), 740-761.
- [9] G. M. Cordeiro, L. H. de Santana, E. M. Ortega, R. R. Pescim, A new family of distributions: Libby-Novick beta. *Int. J. Stat. Probab.* 3(2) (2014), 63.
- [10] M. Ali Ahmed, The new form Libby-Novick distribution, *Communications in Statistics - Theory and Methods.* 50 (2021), 540.
- [11] M. Rashid, Z. Iqbal, M. Hanif, Characterizations and entropy measures of the libby-novick generalized beta distribution. *Adv. Appl. Stat.* 63(2) (2020), 235-259
- [12] Z. Iqbal, M. Rashid, M. Hanif, Properties of the Libby-Novick generalized beta distribution with application. *Int. J. Anal. Appl.* 19 (2021), 360-388.
- [13] G. M. Cordeiro, R. dos Santos Brito, The beta power distribution. *Brazil. J. Probab. Stat.* 26(1) (2012), 88-112.
- [14] R. Dasgupta, On the distribution of burr with applications. *Sankhya B,* 73 (2011), 1-19.